

# Modelling extreme values of processes observed at irregular time step

## Extreme Environmental Events

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# Introduction

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- Spatio-temporal peaks-over-thresholds
- Application: Satellite data, in-situ observations, hindcast data...

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- Univariate extreme values modelling
- Modelling dependence

## 2 Statistical inference

- Theoretical results
- Simulation results

## 3 Conclusion

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# Univariate probabilistic theory

- Limit theorems for  $\frac{\sum_{i=1}^n X_i - b_n}{a_n}$ : CLT,  $\alpha$ -stable distributions...
- For the tails:  $S_n \leftrightarrow M_n = \max_{i=1,\dots,n} X_i$
- If  $F^n(a_n x + b_n) = \mathbb{P}\left\{\frac{M_n - b_n}{a_n} \leq x\right\} \rightarrow G(x)$ , then  $G$  is a max-stable distribution:  $G^n(\alpha_n x + \beta_n) = G(x)$
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$$G(x; \mu, \sigma, \xi) = \exp\left[-\left(1 + \xi \frac{x-\mu}{\sigma}\right)_+^{-1/\xi}\right] \in \text{GEV}(\mu, \sigma, \xi)$$
- Example:
  - Type Weibull:  $\xi < 0$  (including Rayleigh...)
  - Heavy tails (Fréchet):  $\xi > 0$ ; Unit Fréchet  $\xi = 1$
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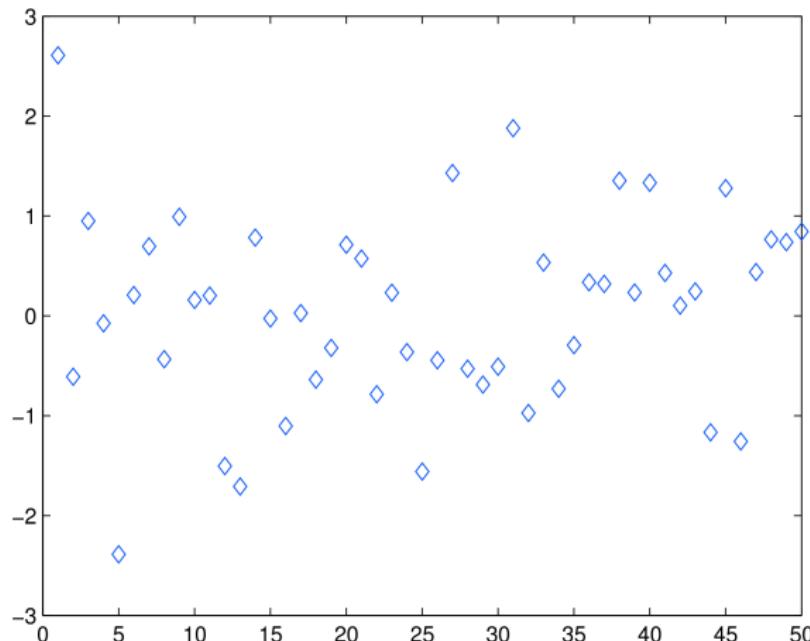
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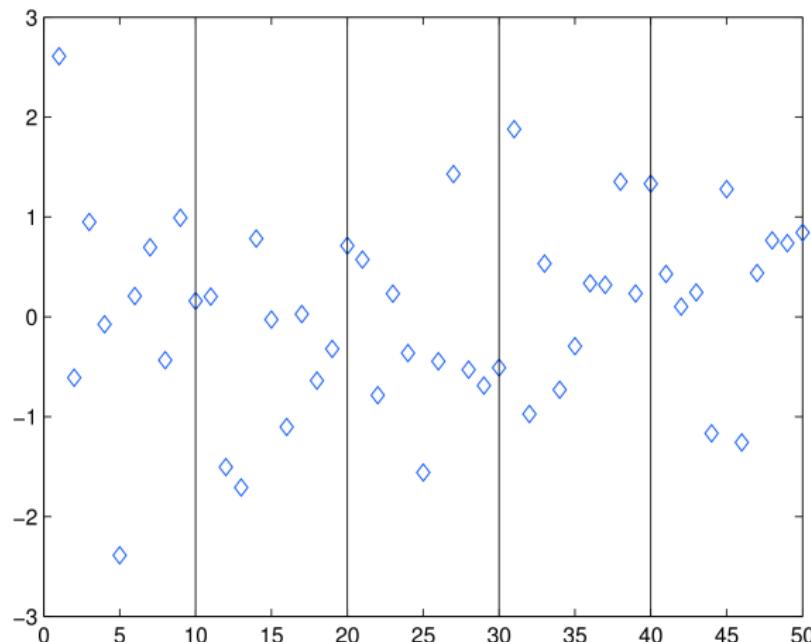
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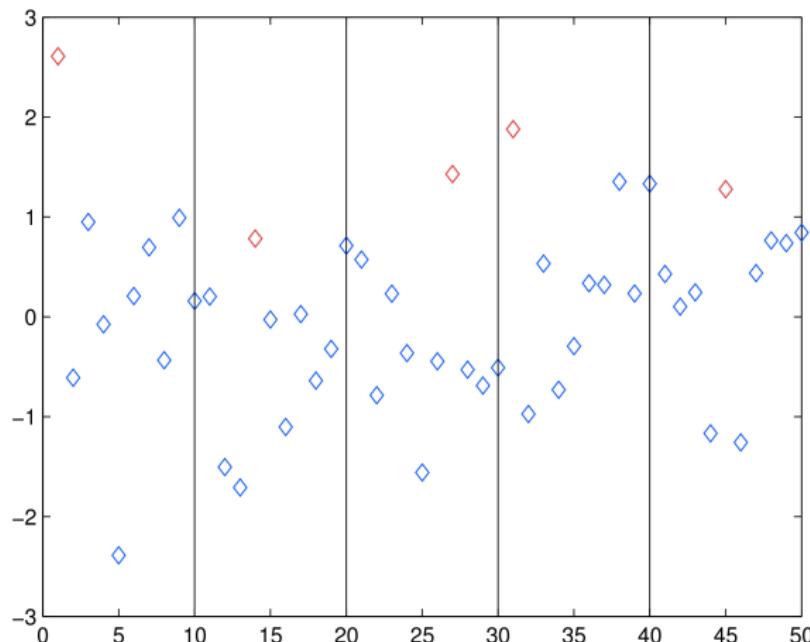
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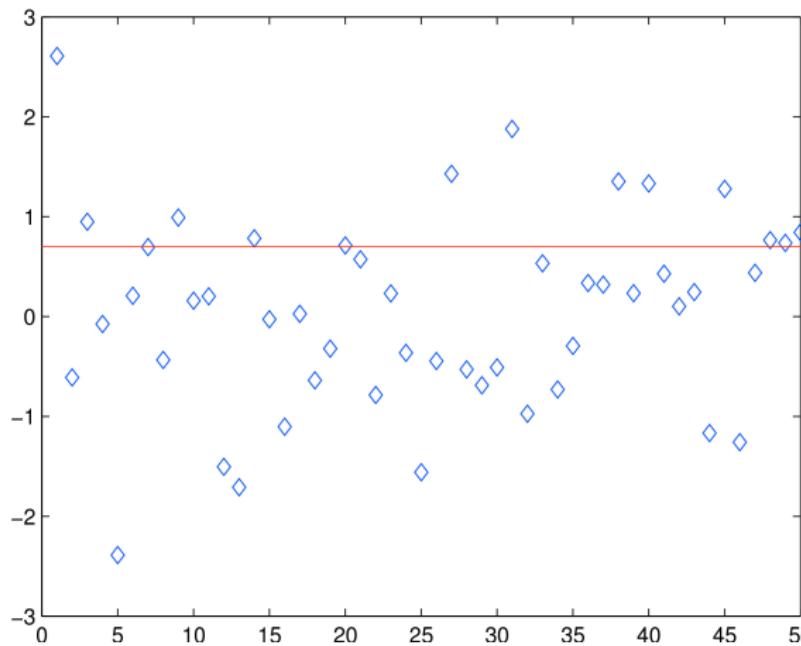
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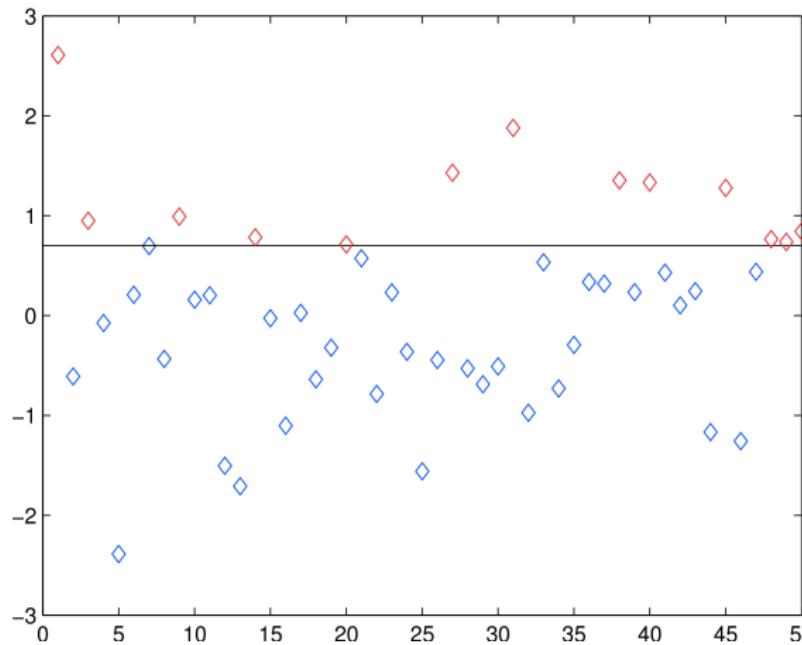
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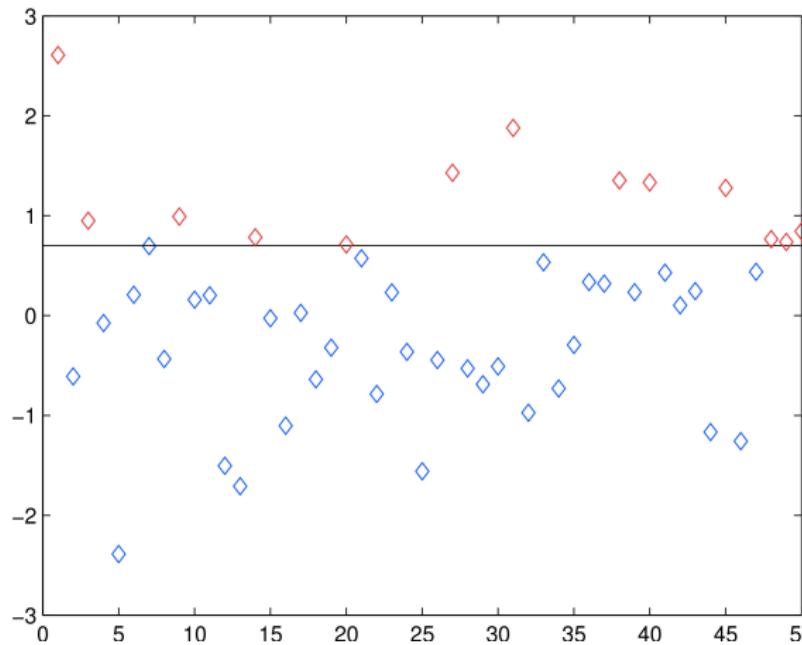
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# Solutions to time-dependence in POT

2 approaches:

## ① Declustering (Coles, 2001)

- Identify Clusters
- Fit the marginal distribution to maximum excess within each cluster
- Estimate the extremal index to **correct** the return period

## ② Modelling all exceedances of the threshold (Smith, 1997; Ribatet 2009)

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# Modelling dependence

## Smith's max-stable process

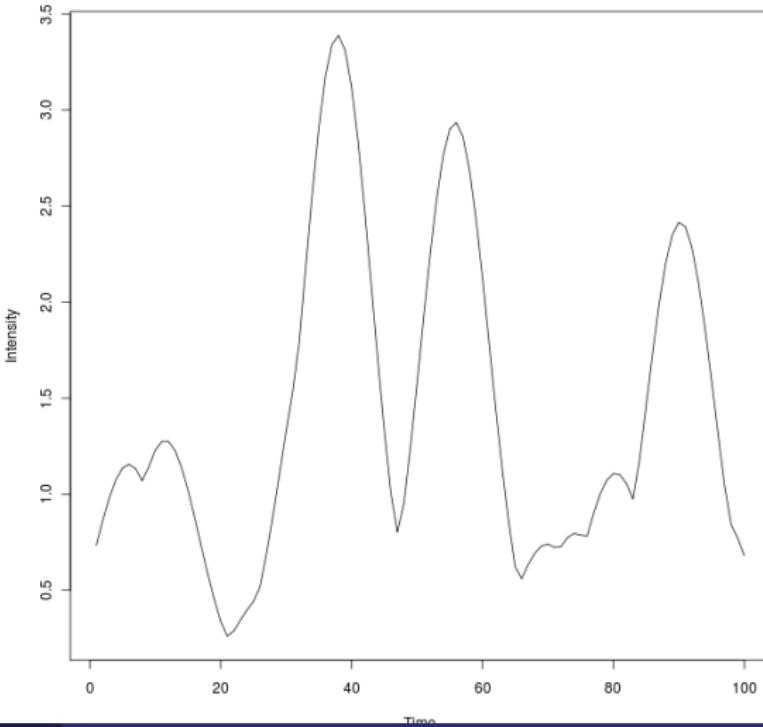
Extension of **max-stable concept for processes** (de Haan, 1986; Smith, 1990), with unit Fréchet marginal distributions.

$$\forall t \in \mathbb{R}, Z(t) = \max_{i=1,\dots} \{\lambda_i f(t, Y_i)\}$$

with:  $f(s, t) = \varphi(s - t)$ ,  $\varphi$  p.d.f of  $\mathcal{N}(0, 1)$ ,  $(\lambda_i, Y_i)$  points of a Poisson process with intensity  $\lambda^{-2} d\lambda/ds$

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$\Rightarrow$  use of a **composite likelihood method** (Varin, 2008):

$$CL(z_{t_1}, \dots, z_{t_n}; \theta) = \prod_{i=1}^{n-K} \prod_{j=i+1}^{i+K} p(z_{t_i}, z_{t_j}; \theta)$$

# Application to peaks over a high threshold

## Model

### Model for excess

- Observations  $\{X_{t_i}\}_{i=1,\dots,n}$
- If  $X_{t_i} > u$ ,  $X_{t_i} = \tilde{X}_{t_i}$  where  $\tilde{X}_t$  is a max-stable process:
  - $\Rightarrow \tilde{X}_t \sim G(\cdot; \mu, \sigma, \xi) \in GEV(\mu, \sigma, \xi)$
  - $\Rightarrow Z_t = -1/\log G(X_t; \mu, \sigma, \xi) \sim \text{Fr\'echet}(1)$
  - $\Rightarrow \{Z_t\}$  is a Smith process
- Thus, for  $x_i > u$  and  $x_j > u$ :

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- Under study: consistency of the uncensored one-lagged CL estimator (i.e. with  $K = 1$ )
- Perspectives:
  - Obtaining asymptotic normality
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  - Choice of  $K$ ?
  - Results with censoring
- Can the behaviour be guessed by simulation?
  - Convergence
  - Influence of parameters values
  - Influence of  $K$
  - Robustness to missing values and censoring

# Properties of the CL estimator

## Properties

- The Smith process is ergodic (Stoev, 2008)
- Under study: consistency of the uncensored one-lagged CL estimator (i.e. with  $K = 1$ )
- Perspectives:
  - Obtaining asymptotic normality
  - Choice of  $K$ ?
  - Results with censoring
- Can the behaviour be guessed by simulation?
  - **Convergence**
  - Influence of parameters values
  - Influence of  $K$
  - Robustness to missing values and censoring

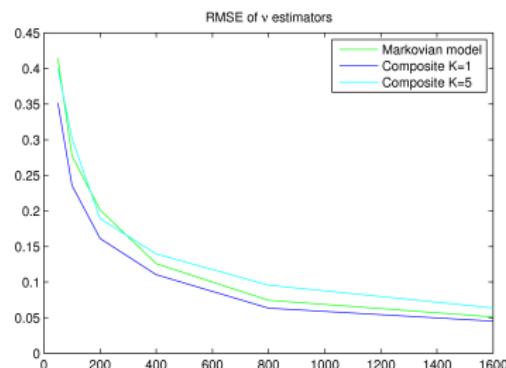
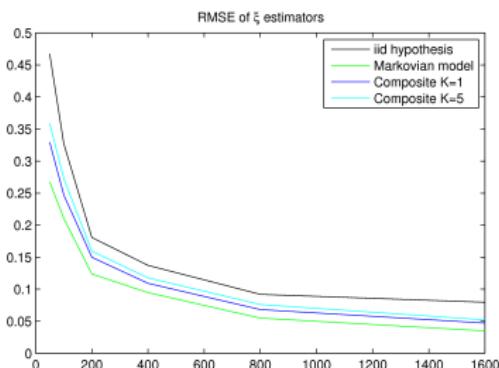
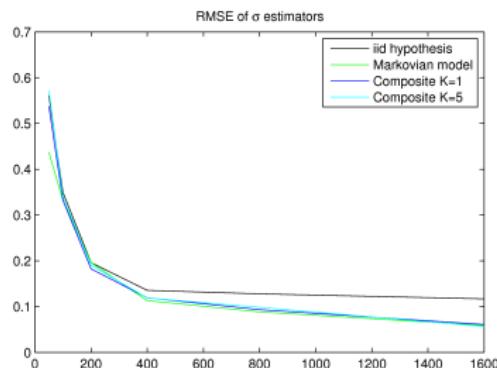
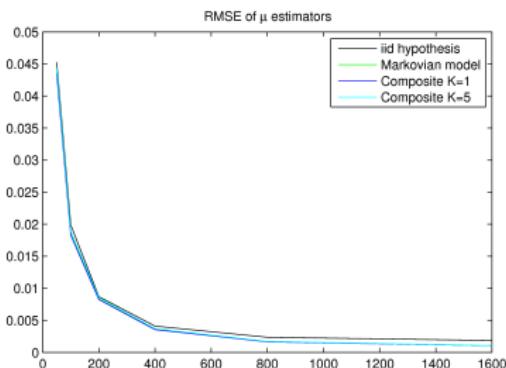


Figure: RMSE of the different estimators

# Robustness

**Question:** Is this model able to catch extremal properties of usual time series models?

- 10 years return level: level which is exceeded on average once every 10 years
- Frequency and length of extremal events
- Sum of the values above the threshold

# Robustness

**Question:** Is this model able to catch extremal properties of usual time series models?

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Method:

- Simulation of  $X_t = \varphi X_{t-1} + \mathcal{N}\left(0, \sqrt{1 - \varphi^2}\right)$  of length 100 years (with 1 obs./day)
- Fitting the model
- 100 times:
  - Simulation of the fitted model over 10 years, and computation of the statistics
  - Simulation of the  $AR(1)$  over 10 years, and computation of the statistics

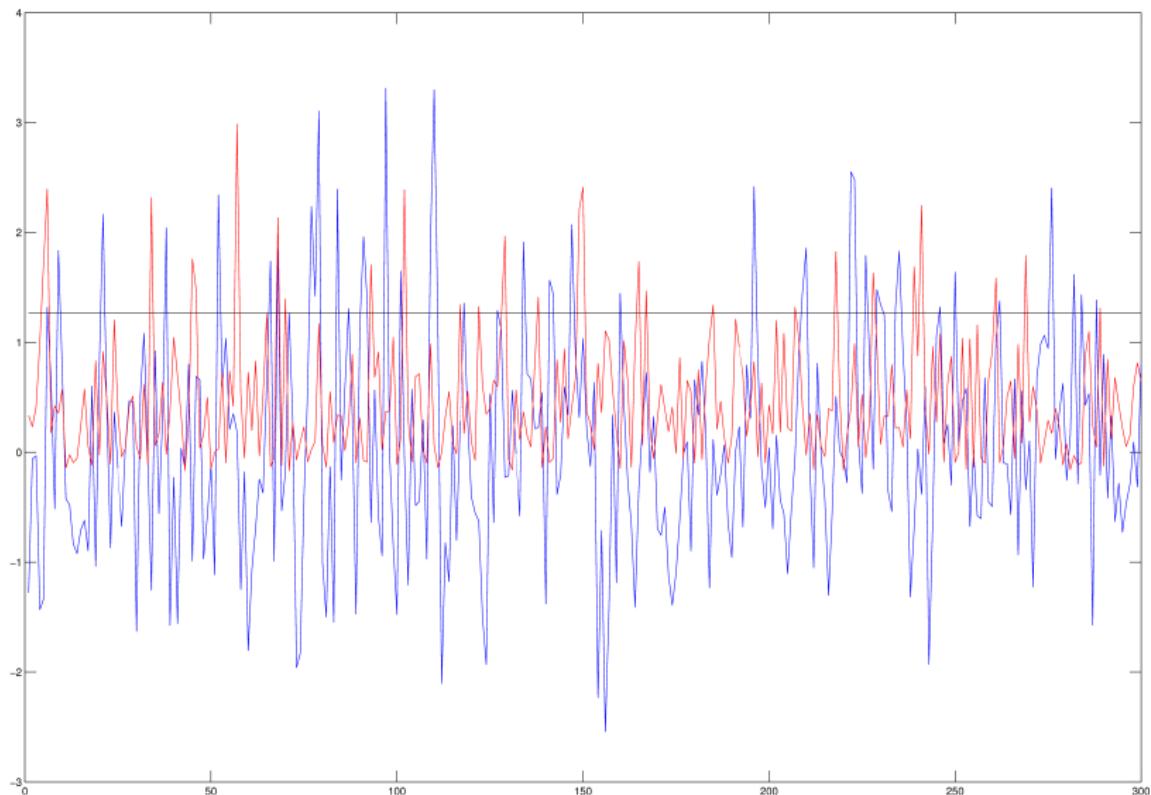


Figure: Example of simulated data (Blue) vs. fitted model (Red) and selected threshold (Black)

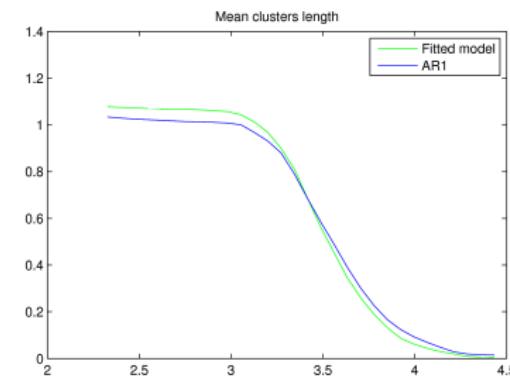
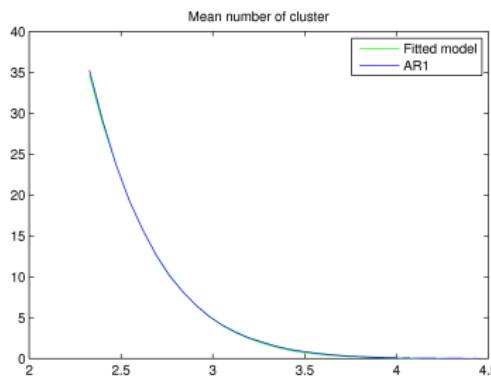
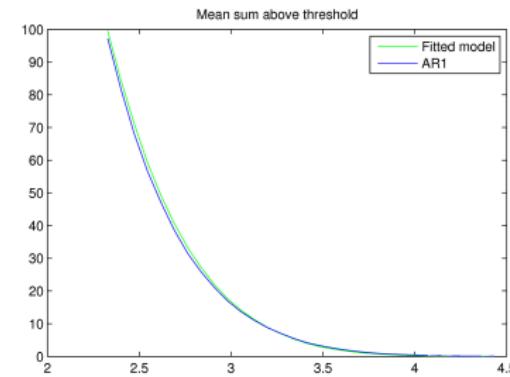
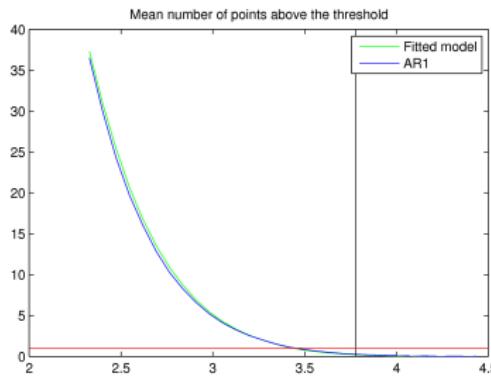


Figure: Statistics computed for an AR(1) process with  $\varphi = 0.2$

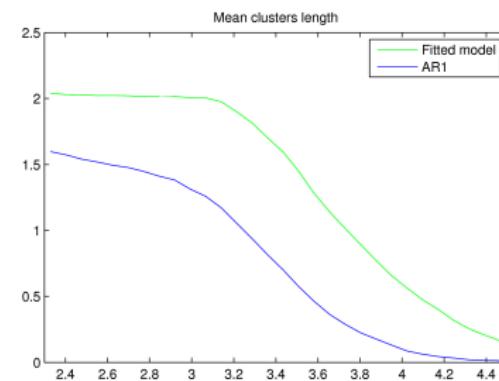
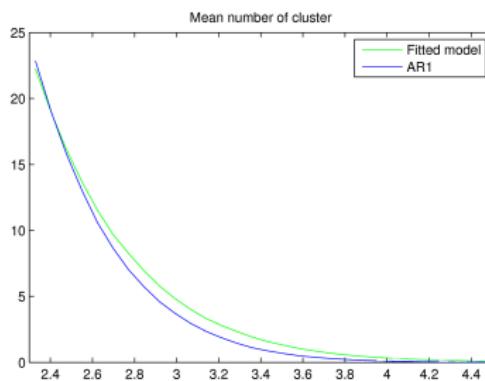
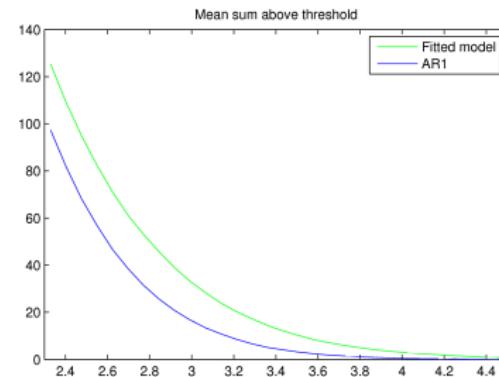
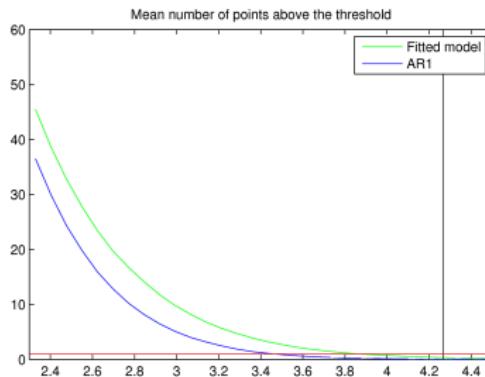


Figure: Statistics computed for an AR(1) process with  $\varphi = 0.8$

# Contents

## 1 Statistical modelling

- Univariate extreme values modelling
- Modelling dependence

## 2 Statistical inference

- Theoretical results
- Simulation results

## 3 Conclusion

# Conclusion

... and prospects

- We obtained a flexible model that can account for irregular data spacing and missing values
- Ongoing: application to satellite and in-situ data
- Perspectives:
  - Theoretical results :asymptotic normality, choice of K, influence of censoring...
  - Extensions: space-time modelling...



## D. R. Cox.

A note on pseudolikelihood constructed from marginal densities.

*Biometrika*, 92:729–737, 2004.



## L. de Haan.

A spectral representation for max-stable processes.

*Ann. Probab.*, 12(4):1194–1204, 1984.



## S. A. Padoan, M. Ribatet, and S.A. Sisson.

Likelihood-based inference for max-stable processes.

*In review*, 2009.



## M. Ribatet, T. B. M. J. Ouarda, E. Sauquet, and J. M. Gresillon.

Modeling all exceedances above a threshold using an extremal dependence structure: Inferences on several flood characteristics.

*Water Resources Research*, 45(3), March 2009.

-  Richard L. Smith, Jonathan A. Tawn, and Stuart G. Coles.  
Markov chain models for threshold exceedances.  
*Biometrika*, 84(2):249–268, 1997.
-  R.L. Smith.  
Max-stable processes and spatial extremes.  
1990.
-  Cristiano Varin.  
On composite marginal likelihoods.  
*AStA Adv. Stat. Anal.*, 92(1):1–28, 2008.